Transmission of *Chlamydia trachomatis* through sexual partnerships: a comparison between three individual-based models and empirical data

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*Chlamydia trachomatis* is the most common bacterial sexually transmitted infection (STI) in many developed countries. The highest prevalence rates are found among young adults who have frequent partner change rates. Three published individual-based models have incorporated a detailed description of age-specific sexual behaviour in order to quantify the transmission of *C. trachomatis* in the population and to assess the impact of screening interventions. Owing to varying assumptions about sexual partnership formation and dissolution and the great uncertainty about critical parameters, such models show conflicting results about the impact of preventive interventions. Here, we perform a detailed evaluation of these models by comparing the partnership formation and dissolution dynamics with data from Natsal 2000, a population-based probability sample survey of sexual attitudes and lifestyles in Britain. The data also allow us to describe the dispersion of *C. trachomatis* infections as a function of sexual behaviour, using the Gini coefficient. We suggest that the Gini coefficient is a useful measure for calibrating infectious disease models that include risk structure and highlight the need to estimate this measure for other STIs.

Keywords: *Chlamydia trachomatis*; sexual partnerships; individual-based model; Natsal 2000; Gini coefficient

1. INTRODUCTION

*Chlamydia trachomatis* is the most common curable sexually transmitted infection (STI) in many developed countries [1]. It is most prevalent among sexually active young adults, with the majority of infections being asymptomatic and remaining undiagnosed [2]. In women, infection can ascend to the upper genital tract, causing pelvic inflammatory disease, which can then lead to tubal factor infertility, ectopic pregnancy and chronic pelvic pain [3]. It is therefore essential to understand the transmission dynamics of *C. trachomatis* in order to develop appropriate interventions aimed at reducing the incidence and prevalence of infections.

Mathematical and computational models of infectious disease dynamics are widely used to study transmission and predict the impact of public health interventions [4–6]. With respect to *C. trachomatis*, such models can help us to understand the potential impact of screening, a population-level intervention that aims to reduce chlamydia prevalence and associated complications through the regular detection and treatment of asymptomatically infected cases. There are no empirical studies that show what levels of coverage and frequency of screening are necessary to achieve a specified reduction in *C. trachomatis* prevalence. The predictions from modelling studies can therefore be influential in guiding health policy decisions about chlamydia screening [7].

The models that have been developed to investigate the impact of screening have used either deterministic, population-based [8–10] or stochastic, individual-based approaches [11–13]. The latter approach is...
sometimes favoured, because it allows the detailed description of sexual partnership dynamics. Taking sexual partnerships explicitly into account has been suggested as an important feature of models of chlamydia transmission to capture re-infection within partnerships and investigate the effect of partner notification [10,14]. In spite of these advantages, it is challenging to appropriately parameterize individual-based models and analysis and interpretation of the results is generally difficult [15,16]. As a matter of fact, the three individual-based models of C. trachomatis transmission, which were designed to evaluate different screening interventions in the UK and The Netherlands, can show marked differences in the predictions which complicates their use in informing public health policy-making. The study by Kretzschmar et al. [17] described key features of each model and compared the effect of implementing the same screening scenario in the three models. The predicted reductions in chlamydia prevalence in women after 10 years of the intervention were 85 per cent [12], 25 per cent [11] and 5 per cent [13].

The contrasting predictions between the three individual-based models are a result of different assumptions about the sexual partnership dynamics and the infection parameters of C. trachomatis. These are key determinants that are believed to affect the spread of STIs. Chlamydia trachomatis, like many other STIs, is primarily found among sexually active young adults who change sexual partners frequently. This is a prerequisite for sustaining transmission of bacterial STIs with infectious periods between several months to years. Further key determinants are the duration of sexual partnerships, the gap interval between subsequent partnerships, concurrent partnerships and heterogeneity in sexual behaviour [18–20]. In order to model the spread of C. trachomatis within the population, and to obtain reasonable predictions of screening interventions, a solid comprehension of these behavioural characteristics is necessary.

Novel insights into the dynamics of sexual contacts have been gained during the last decade [21–23] but no clear formalism to describe the individual’s sexual behaviour has been established so far. Also, the necessary parameters that define duration of partnerships, concurrency or age-specific behaviour are exceedingly difficult to obtain and implement in mathematical and computational models. As an additional complication, many of these behavioural characteristics are generated by a multitude of parameters. In order to assess the ‘realism’ of such models, it is therefore necessary to evaluate them in light of empirical data. On the one hand, this allows decisions about which models and predictions should be favoured over others. On the other hand, the features and structure of the most ‘realistic’ models can also give insights into the underlying mechanisms of sexual partnership formation. One of the most comprehensive data sets available to date is Natsal 2000, a population-based probability sample survey of sexual attitudes and lifestyles in Britain [24]. The large sample size of the survey allows the derivation of the characteristic distributions of measures of sexual behaviour in a population. In addition, C. trachomatis prevalence was measured in a subset of sexually active respondents [19]. Natsal 2000 therefore provides a unique opportunity to study certain characteristics of sexual behaviour and link them to the prevalence of C. trachomatis.

In this paper, we aim to explain the key differences of the three individual-based models that resulted in contrasting predictions of the impact of screening interventions against C. trachomatis [11–13,17]. By comparing the sexual partnership dynamics of the models to data from Natsal 2000, we find several properties that can affect the transmission of C. trachomatis. We illustrate the various assumptions made when describing sexual partnerships and discuss their implications for the spread of C. trachomatis and the impact of a standardized screening intervention. This allows us to report the most reasonable prediction from a specific screening intervention. As a useful summary measure to evaluate models of STI transmission, we propose that the distribution of infections among individuals with different levels of sexual activity should be described using the Gini coefficient.

2. METHODS

2.1. Individual-based models

A detailed comparison of key features of the three models has been presented previously [17].

2.1.1. Kretzschmar model. The Kretzschmar model was developed at the Dutch National Institute of Public Health and the Environment, to evaluate strategies for controlling C. trachomatis and Neisseria gonorrhoeae [11]. The simulated population in the model consists of 15–64 year olds who belong to one of the two sexual activity groups; a small core group (5%) of younger adults (<35 years) who, in addition to steady relationships, undergo casual relationships that can be concurrent (with a maximum of two partners at any given time), and the other group who are either in a steady or (less likely) in a casual relationship.

2.1.2. Turner model. The Turner model was originally developed at the Health Protection Agency (HPA), London, UK [12], based on an earlier model for N. gonorrhoeae transmission developed by Ghani et al. [25]. The model simulates heterosexual partnership formation between 16 and 44 year olds. Individuals in the population have a preference for either short or longer duration partnerships. Additionally, a small core group (5%) can undergo a concurrent partnership as long as they are younger than 35 years.

2.1.3. ClaSS model. The Chlamydia Screening Studies (ClaSS) model was developed in the UK as part of the ClaSS project [13]. The simulated population consists of 12–65 year olds who undergo heterosexual partnership formation. Individuals’ sexual behaviour is divided into three activity classes with all people older than 35 year moving into the lowest activity class. The partnership duration and the possibility of forming concurrent partnerships (with up to eight concurrent partners) depend on the activity classes of both individuals. For
Table 1. Parameters describing infection with C. trachomatis. Assuming the same average duration of the infection ensures that the baseline turnover of C. trachomatis is the same in all three individual-based models. However, we adjust the transmission rate per sex act to obtain a similar C. trachomatis prevalence. Analysis of data from Natsal 2000 results in a prevalence of 1.7% (95% CI: 1.3–2.1) for people aged 18–44 years and 2.9% (95% CI: 1.3–2.1) for people aged 18–29 years [19]. For a detailed evaluation of disease-specific parameters, see Althaus et al. [9]. Note that the Turner model does not include an incubation time.

<table>
<thead>
<tr>
<th>Parameters</th>
<th>Value</th>
</tr>
</thead>
<tbody>
<tr>
<td>incubation stages</td>
<td></td>
</tr>
<tr>
<td>incubation time, i.e. the time people are infected but not yet infectious</td>
<td>14 days</td>
</tr>
<tr>
<td>fraction of infections becoming asymptomatic</td>
<td>0.75</td>
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<tr>
<td>duration of symptomatic period</td>
<td>35 days</td>
</tr>
<tr>
<td>duration of asymptomatic period</td>
<td>433 days</td>
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<tr>
<td>transmission rate per sex act</td>
<td></td>
</tr>
<tr>
<td>Kretzschmar model</td>
<td>4.4%</td>
</tr>
<tr>
<td>Turner model</td>
<td>3.5%</td>
</tr>
<tr>
<td>ClaSS model</td>
<td>2.2%</td>
</tr>
<tr>
<td>average interval between sex acts</td>
<td></td>
</tr>
<tr>
<td>during the first two weeks of a partnership</td>
<td>2 days</td>
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<tr>
<td>after the first two weeks of a partnership</td>
<td>7 days</td>
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In this study, the ClaSS model was implemented in the package Rtsisim (manuscript in preparation) of the R software environment for statistical computing [26].

2.1.4. Chlamydia trachomatis transmission and screening intervention. We simulate transmission of C. trachomatis in the three individual-based models to study the endemic steady-state prevalence and the impact of a standardized intervention programme. We have shown recently that different assumptions about the transmission dynamics and turnover of C. trachomatis can affect the predicted impact of a screening programme [9]. Therefore, we use a harmonized parameter set to describe the duration of infection and the proportion of symptomatic and asymptomatic cases (table 1). The values chosen are based on a recent evaluation of parameters that have been used in previous mathematical and computational models [9] and on consensus between the authors. We then adjust the transmission probability per sex act in each model in order to obtain the same baseline prevalence. After running the models such that the partnership dynamics and the baseline prevalence have equilibrated, we introduce a standardized intervention programme to reduce the prevalence of C. trachomatis. Eligible for screening are all sexually active people younger than 25 years. The proportion of people that gets screened at least once a year, p, is 20 per cent. From the relation \( p = 1 - e^{-ct} \), we obtain a screening rate of \( c = 0.223 \) yr\(^{-1}\). Note that some people can receive screening more than once a year. If people have a positive chlamydia test result, 40 per cent of their current or most recent partners will be notified and successfully treated. The predicted impact on the prevalence of C. trachomatis is reported after 5 and 10 years.

2.2. Sexual behaviour data

We use data from Natsal 2000, a population-based probability sample survey of sexual attitudes and lifestyles in Britain [24]. The sample consists of 11 161 women and men aged 16–44 years. Urine samples for ligase chain reaction testing for C. trachomatis are available for a subset of 3569 sexually active respondents aged 18–44 years [19]. Weighted data (to adjust for unequal selection probabilities and to correct for the age and gender profile in the population) were used to derive measures of sexual behaviour and C. trachomatis prevalence. We consider the following measures of sexual partnerships: lifetime number of heterosexual partners; number of heterosexual partners during the last year and the date of first and last heterosexual intercourse with the most recent and second most recent partner. The latter allow us to obtain the distribution of partnership durations and the length of gaps or overlaps between two partnerships. All data were analysed using the R software environment.

2.3. Estimating partnership and gap durations

In cross-sectional surveys, the start of a sexual partnership can, in principle, be clearly defined. However, information about the end of a partnership is more difficult to obtain as reported partnerships might be ongoing. This results in an oversampling of longer partnerships and inflated partnership durations [27–29]. We focus our analysis on the second most recent partnership, which allows us to capture short partnership durations as well. In Natsal 2000, the duration of the second most recent partnership was defined as the time in months between first and last sexual intercourse with the second most recent partner within the last 5 years. We further define the duration of the gap between the most recent and second most recent partnership as the later date of first sexual intercourse in either partnership minus the date of last sexual intercourse with the second most recent partner. Thus, we obtain the gap durations for three different scenarios: the first scenario is when the second most recent partnership begins, resulting in a positive gap duration (figure 1a); in the second, the second most recent partnership has...
not ended yet before the most recent partnership begins, resulting in a negative gap duration or overlap (concurrent partnership) (figure 1b); in the third, the second most recent partnership begins and ends during the duration of the most recent partnership, also resulting in an overlap (figure 1c). Since the dates of first and last sexual intercourse are reported in months, a gap duration of zero months can result from a short gap of less than or equal to 1 month or a short overlap of not more than a month.

The analysis of sexual partnership durations can be prone to additional selection biases [30]. However, our primary interest is not to estimate the true duration of partnerships and gaps but rather to compare the simulated data with empirical data. We therefore perform the same cross-sectional survey as Natsal 2000 in the three individual-based models. The following algorithm was applied to the simulated population of 16–44 year olds to obtain the partnership duration and gap length: (1) all current and past partnerships of an individual are sorted by their end date in descending order; (2) if partnerships are ongoing, the end date of the partnership is the time at cross-section; (3) the second most recent partnership is defined as the one with the second most recent end date. However, if more than one partnership is ongoing, one will be randomly selected as the second most recent; (4) the duration of the second most recent partnership is given as the time between the start and end date of the second most recent partnership; and (5) the duration of the gap between the most recent partnership is given as the time between the end date of the most recent partnership and the end date of the second most recent partnership.

### 2.4. Lorenz curve and Gini coefficient

The Lorenz curve is a graphical representation of an empirical cumulative probability distribution. It is most well known as a method for representing inequality in income or wealth [31]. Recently, it has been applied to visualize the geographical distribution of several bacterial STIs at different levels [32–35] and over time [36]. In a mathematical modelling study, it has also been used as a method to describe the distribution of infections across certain subpopulations [37]. Here, we adopt this framework and use the Lorenz curve to represent the asymmetric distribution of *C. trachomatis* infections among individuals with *n* different levels of sexual activity. Thus, the Lorenz curve represents the cumulative proportion of *C. trachomatis* infections (\(y_i\)) as a function of the cumulative proportion of the population (\(x_i\)), after population sub-groups \(i (i = 1, 2, \ldots, n)\) have been ranked in the order of their sexual activity level (i.e. number of partners within the last year). The diagonal line on a Lorenz curve plot denotes perfect equality, i.e. every population sub-group has the same prevalence of *C. trachomatis*. As a summary measure for the asymmetrically distributed infections, we can calculate the Gini coefficient [38]. It is defined as the area between the line of equality and the Lorenz curve over the total area below the line of equality:

\[
G = 1 - \sum_{i=1}^{n} (x_i - x_{i-1})(y_i + y_{i-1}),
\]

where \(x_0 = y_0 = 0\) and \(x_n = y_n = 1\). A Gini coefficient of 0 represents perfect equality, whereas a Gini coefficient close to 1 implies that only a small proportion of the population that exhibits a high level of sexual activity is infected with *C. trachomatis*. We plot the Lorenz curve and estimate the Gini coefficient for the empirical data from Natsal 2000 together with the distribution of *C. trachomatis* infections that we obtain from the three individual-based models.

### 3. RESULTS

#### 3.1. Sexual partnership dynamics

In order to gain insights into differences between the three individual-based models, we compare them to four characteristic patterns of sexual partnership dynamics from Natsal 2000: the lifetime number of partners; the number of partners in the last year; the duration of partnerships; and the duration of gaps between partnerships.

##### 3.1.1. Lifetime number of partners

Empirical data about the distribution of the lifetime number of partners have been found to resemble a power-law, suggesting that sexual-contact networks exhibit features of scale-free networks [21,39]. While there is an ongoing debate as to what extent sexual networks have such an underlying structure [40–42], the primary feature of the distribution of the lifetime number of partners is that there is a small fraction of people with a very high number of partners; more than would be expected if individuals engaged in sexual contacts at random. Although mechanisms that would generate such a pattern have been proposed [22,23], the individual-based models that we evaluate here are based on the classical assumption that individuals in different ‘risk’ groups exhibit a certain degree of assortative mixing. It can be seen that assuming two (Kretzschmar model) or three (ClaSS model) risk groups can approximate the empirical distribution to a certain extent (figure 2). In contrast, the Turner model assumes more random mixing, which results in a Poisson-like degree distribution. All models clearly fail to describe individuals with a very high number of partners to date (>500) which is mainly due to the limited population size in the simulations. However, it is unlikely that those people contribute substantially to the transmission of *C. trachomatis*. Very high numbers of lifetime partners are generally reached when individuals are beyond the age where they significantly contribute to the transmission of *C. trachomatis*.

##### 3.1.2. Number of partners within the last year

An important characteristic of *C. trachomatis* is that most transmission occurs among sexually active young adults. One of the aims of individual-based models is to account for age-specific sexual behaviour. Empirical
data from Natsal 2000 show an increase in the average number of heterosexual partners within the last year for people between the ages of 16 and 19 years (figure 3). The partner change rate decreases thereafter and stabilizes after the age of 35 at an average of around one partner per year. The Kretzschmar model does not explicitly include age-specific partner change rates, but instead uses a relatively simple algorithm where a small fraction of individuals (5%) belongs to a core group with increased sexual activity. At the age of 35 years, everyone moves into the low-risk population. The ClaSS model incorporates detailed age-specific sexual behaviour which follows a similar pattern to that shown in the empirical data. However, it underestimates the average number of partners, especially for individuals older than 35 years. Although the Turner model captures the partner change rates of individuals older than 20 years relatively well, it greatly overestimates the sexual contact rates for people of age 16–19 years. This feature of the model has been noted previously and results in an overestimation of *C. trachomatis* transmission among young people [17].

### 3.1.3. Duration of sexual partnerships

The duration of sexual partnerships is critically important for the transmission of *C. trachomatis*. Individuals with frequent partner change rates are expected to have many short-term (casual) partnerships. On the other hand, the probability of *C. trachomatis* transmission increases in long-term (steady) partnerships because of the higher number of sex acts. The cumulative distribution of sexual partnership durations in Natsal 2000 shows a biphasic pattern with a large proportion of short partnerships and exhibits a long tail for longer partnerships (black solid line). Individual-based models generally assume exponential distributions of partnership duration and only mimic the empirical data if they explicitly account for short (casual) partnerships, such as the Kretzschmar model (red dashed line) and the Turner model (green dotted line). The blue dashed-dotted line denotes the ClaSS model. The graph represents data from the second most recent heterosexual partnership of all women and men aged 16–44 years. (Online version in colour.)
All three individual-based models that
ships, respectively, has been found to be a particularly
and overlaps in serial and concurrent sexual partner-

3.1.4. Duration of gaps and overlaps.

Kretzschmar and Turner models distinguish between
casual and steady partnerships and can approximate
the empirical data to a certain extent. In contrast, the
distribution of partnerships in the ClaSS model appears
to follow a single exponential distribution, underestimat-
ing the proportion of casual and steady partnerships.
The continuously decreasing slope in the distribution of
partnership durations with ongoing time indicates a
certain degree of robustness. This property could be
taken into account in individual-based models by setting
the probability that a partnership dissolves as a function
of the duration of the partnership.

3.2. Dispersion of C. trachomatis infections

Given that the sexual partnership dynamics between
the three individual-based models can vary widely, it is
natural to investigate how C. trachomatis spreads in
the simulated populations. In previous studies, the
three models have used different assumptions about
the duration of infection and the proportion of sympto-
matic and asymptomatic cases [17]. Here, we use a
harmonized set of C. trachomatis infection parameters,
which is based on a thorough evaluation of the par-
parameter ranges in the literature and a recent estimate
of the duration of the asymptomatic period (table 1).
By adjusting the transmission rate per sex act in the
three individual-based models, we can roughly repro-
duce the estimated prevalence of C. trachomatis from
Natsal 2000 (1.7% (95% CI: 1.3–2.1) for people aged
18–44 years and 2.9% (95% CI: 1.3–2.1) for people
aged 18–29 years) [19].

The C. trachomatis prevalence data in Natsal 2000
are unique, because we have detailed information on
the sexual behaviour of each individual. This gives us
the opportunity to study the dispersion of C. trachoma-
tis infections across the population and the fraction of
infections that is concentrated in individuals that exhi-
bbit frequent partner change rates. Not unexpectedly,
the prevalence of C. trachomatis increases with an
increasing number of heterosexual partners within the
last year (figure 6a). This corroborates the concept of
a ‘core group’, illustrated in the seminal work about
N. gonorrhoeae by Hethcote & Yorke [47]. However,
C. trachomatis seems to be less concentrated in ‘high-
risk’ individuals compared with N. gonorrhoeae with
48 per cent of all chlamydia infections being among
individuals with not more than one partner during
the last year. Comparing the data from Natsal 2000
with the distribution of infections in the three individ-
ual-based models provides important information. The
Kretzschmar model captures the overall distribution
reasonably well and provides a striking resemblance of
the proportion of people in each ‘risk category’ (com-
pare the legend of figure 6a,b). The prevalence in the
group of individuals with five or more partners within
the last year is somewhat overestimated, however. The
Turner model also captures the overall distribution of
C. trachomatis infections reasonably well, with the
exception that no individual remains without a partner
during the period of 1 year. However, we have shown
above that the Turner model assumes an unusually
high partner change rate for young adults aged 16–19
years (figure 3). This behaviour is not captured in
figure 6c as we only consider individuals aged 18–44
years, reflecting the age group for which C. trachomatis
positivity is available in Natsal 2000. In contrast to the

Figure 5. Cumulative distribution of the gap length. Negative
gaps denote overlaps between partnerships and positive gaps
denote waiting times until new partnerships are formed.
Owing to different assumptions of the partnership formation
process, individual-based models exhibit different gaps and
overlaps. The graph represents data about gaps or overlaps
between the most recent and the second most recent hetero-

sexual partnership from all women and men aged 16–44
years. The black solid line denotes Natsal 2000; the red
dashed line denotes the Kretzschmar model; the green
dotted line denotes the Turner model and the blue dashed-
dotted line denotes the ClaSS model. (Online version in
colour.)
other two models, \textit{C. trachomatis} infections are too heavily concentrated among ‘high-risk’ individuals in the ClaSS model (figure 6d). Further, it becomes apparent that the sexual partnership dynamics are not captured well as too many individuals remain single during the period of 1 year, corroborating the long duration of gaps in this model.

Given the empirical and simulated data from figure 6, we can plot the Lorenz curve that represents the cumulative proportion of \textit{C. trachomatis} infections as a function of the cumulative proportion of the population, after population sub-groups have been ranked in the order of their sexual activity during the last year (figure 7). From the Lorenz curve, we can calculate the Gini coefficient, which gives us a statistical measure for the strength of asymmetry in the distribution of \textit{C. trachomatis} (see §2). The Kretzschmar (0.46) and the Turner models (0.32) are close to the estimated Gini coefficient from Natsal 2000 (0.38), whereas the ClaSS model has a much higher Gini coefficient (0.84).

### 3.3. Impact of screening interventions

Given the different dispersion of \textit{C. trachomatis} infections in the three models, we aim to explain the different results that these models show upon introduction of a standardized screening programme [17]. We implement the same standardized intervention (see §2) and compare the predicted impact after 5 and 10 years of intervention (figure 8). In the previous study, the Turner model predicted the strongest impact of an intervention, whereas the ClaSS model showed almost no impact at all [17]. Since we use a harmonized disease-specific parameter set between the models, we now find more similar results for the Kretzschmar and the ClaSS models. However, the Turner model still predicts a much stronger impact of the same intervention. One reason for this discrepancy in the Turner model is the overestimation of partner change rates among young adults (figure 3). The unrealistically high partner change rates imply that a large proportion of \textit{C. trachomatis} transmission occurs in exactly those individuals who become eligible for screening, rendering the intervention more powerful. A second reason could be the short duration of gaps between partnerships, which results in very few individuals remaining single. Thus, infected individuals are usually engaged in a partnership with the possibility of transmitting the infection. Screening and treatment of an index case will therefore result in a reduction of \textit{C. trachomatis} transmission at the population level. In contrast to the Turner model, the ClaSS model overestimates the proportion of individuals that are single.
Chlamydia trachomatis infections in those individuals are likely to clear naturally before they engage in a new partnership. Thus, screening and treatment of infected singles in the ClaSS model has less effect as it does not necessarily reduce transmission at the population level.

We can also express the reduction in the prevalence of C. trachomatis as a percentage of the pre-intervention steady state. After 10 years of screening, we obtain the same order for the group of 16–24 year olds (ClaSS: 39%, Kretzschmar: 26%, Turner: 3%) and the 16–44 year olds (ClaSS: 49%, Kretzschmar: 34%, Turner: 7%). Interestingly, we find the same order for the Gini coefficients obtained in §3.2 and hypothesize that these coefficients are one determinant of the predicted impact of an organized screening programme. A high Gini coefficient suggests that most C. trachomatis infections are concentrated in a small fraction of the population that exhibits high partner change rates. Thus, after infected individuals are treated for C. trachomatis, they rapidly engage in new sexual partnerships where they can contract the infection again, diminishing the effect of a screening programme. In contrast, a low Gini coefficient derives from a more equal dispersion of C. trachomatis infections among individuals with different numbers of partners. A large proportion of all infections is concentrated among individuals with low partner change rates. These individuals are less likely to contract the infection after treatment and partner notification becomes increasingly efficient as it reduces the risk of re-infection in an ongoing relationship.

4. DISCUSSION

Our evaluation of the three individual-based models allowed us to find key reasons that affect the transmission of C. trachomatis and result in contrasting predictions of the impact of screening interventions. Different assumptions about the individual’s sexual behaviour give rise to a widely different pattern of the sexual partnership dynamics between the models. This illustrates that describing some critical characteristics of the sexual partnership dynamics between individuals remains a challenging goal. Although all models capture some aspects reasonably well, they can fail in others. Overall, we find that the Kretzschmar model performs best in describing the sexual partnership dynamics as well as the transmission of C. trachomatis. This is striking, because it was the first model of its kind and its description is based on relatively simple algorithms and a moderate number of parameters. This shows that the most essential characteristics of the sexual partnership dynamics at the population level can be reasonably captured if simulated individuals are subjected to a small set of rules. It also emphasizes the fact that care should be taken if one wants to increase the complexity of models in order to achieve more ‘realism’. As several processes of the sexual partnership dynamics and their corresponding parameters at the individual level are not well understood, such attempts can result in models with ‘unrealistic’ outcomes. A critical point is that sexual behaviour surveys do not necessarily provide the data to parameterize single individuals in a model. We therefore suggest a careful comparison of the sexual partnership dynamics that emerge from individuals’ behaviours in individual-based models, with the corresponding population-based data from sexual behaviour surveys.
The models that we have evaluated here are based on the classical framework where individuals are divided into subpopulations (compartments) such as risk groups that exhibit characteristic properties. This classification is often made arbitrarily and does not allow the individual’s heterogeneity to be fully taken into account. The individual-based paradigm, however, allows this framework to be extended. Individuals do not need to belong to a certain subpopulation or group that will define their behaviour. Instead, the behaviour of an individual can emerge from a simple set of rules that takes into account their own and their partner’s history. De Blasio et al. [23] have shown that the previous sexual history of individuals determines their future sexual contacts to a certain degree. It will be interesting to apply such mechanisms to STI modelling as they can account for a wider variation in sexual behaviour, compared to classical approaches.

The last decade has shown an increasing trend towards sophisticated and detailed computational models to describe the transmission of various infectious diseases and to assess the potential impact of public health interventions [15]. Often, different models are being built for the same purpose, which makes comparisons important to assess the reliability of their results and predictions. For example, the increasing number of models describing influenza pandemics has resulted in some formal comparison studies [48,49]. Such comparison studies can help to understand the impact of key assumptions in the models.

As we have shown here, it is also necessary to re-evaluate models in light of more recent and more detailed data. The present study advances the findings of our previous comparison of the individual-based models of *C. trachomatis* transmission [17]. By harmonizing the disease-specific parameters, the differences in the predicted impact of screening are reduced.

In addition, the transmission of *C. trachomatis* is affected by the marked differences in the sexual partnership dynamics that we find between the models. These differences can at least partly explain the differences in the outcomes between the models that we find here and that have been previously described [17]. Overall, we find that the Kretzschmar model generally performs best in terms of describing the sexual partnership dynamics and the dispersion of *C. trachomatis* infections within the population. We therefore conclude that its predictions are likely to be the most reasonable. The predicted reductions in prevalence in the Kretzschmar model are comparable to those from simpler deterministic models of *C. trachomatis* transmission that examine similar screening interventions [8–10]. This is an important finding, which shows that many questions regarding the impact of public health interventions against *C. trachomatis* can be reasonably investigated with simpler mathematical models [50].

In terms of public health policy-making, the results suggest that screening all sexually active young adults at intervals of 2–5 years (corresponding to a yearly coverage of about 20–40%) for 5–10 years could potentially reduce the prevalence of *C. trachomatis* substantially. Unfortunately, there are still no empirical data from randomized controlled trials of chlamydia screening interventions to show that this kind of reduction in chlamydia prevalence is achievable in practice.

Finally, we have introduced a new summary measure to describe the distribution of *C. trachomatis* infections across the population as a function of people’s sexual behaviour. By combining epidemiological data with behavioural data, this measure, which can be expressed as a Gini coefficient, provides insight into both the sexual partnership dynamics and the transmission of *C. trachomatis*. Owing to the difficulties in parameterizing individual-based models, such summary measures are especially suited to evaluation of the properties of such models. In particular, age- and gender-specific Gini coefficients will provide further insight into the transmission of *C. trachomatis* and public health interventions could be designed accordingly. We suggest that the Gini coefficient is a useful measure for calibrating infectious disease models that include risk structure and highlight the need to estimate this measure for other STIs.

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